

240A Cameron Winter 2006
Department of Economics, U.C.-Davis

Midterm Exam: February 8

Compulsory. Closed book. Worth 32% of course grade.

Read question carefully so you answer the question.

Keep answers as brief as possible.

Question scores (total 50 points)

Question	1a	1b	1c	1d	2a	2b	2c	3a	3b	3c	3d	4a	4b	4c	5a	5b	5c	6a	6b	6c	6d
Points	2	2	2	2	2	2	2	2	2	4	2	4	2	2	2	4	2	2	2	3	3

1. Given sample values of (x, y) equal to $(0, 0)$, $(2, 4)$ and $(4, 2)$, OLS regression of y on x and intercept gives $\hat{\beta}_1 = 1.0$ and $\hat{\beta}_2 = 0.5$. **There is no need to derive these estimates.**

Answer the following giving the appropriate algebra (either summation or matrix notation is okay).

[If your calculator performs regression do not use this feature. Use first principles.]

(a) Show that s , the standard error for this regression, equals 2.4495.

(b) Calculate $s_{\hat{\beta}_2}$, the standard error of $\hat{\beta}_2$. [Use $s = 2.4495$ even if you cannot answer (a)].

(c) Calculate a 95 percent confidence interval for β_2 .

[For t with 1, 2 and 3 degrees of freedom, the critical values are 12.706, 4.303 and 3.182.]

(d) Calculate R^2 for this regression.

2. Suppose \mathbf{b} is a 2×1 vector with entries b_1 and b_2 and \mathbf{A} is a 2×2 matrix with entries a_{11} , a_{12} , a_{21} and a_{22} .

(a) Obtain the formula for $\mathbf{b}'\mathbf{A}\mathbf{b}$.

(b) Given (a), calculate $\partial \mathbf{b}'\mathbf{A}\mathbf{b} / \partial b_1$ and $\partial \mathbf{b}'\mathbf{A}\mathbf{b} / \partial b_2$.

(c) Verify that your answer in (b) implies $\partial \mathbf{b}'\mathbf{A}\mathbf{b} / \partial \mathbf{b} = 2\mathbf{A}\mathbf{b}$ if \mathbf{A} is symmetric..

3. Consider the usual multiple regression model with intercept, which can be written as $\mathbf{y} = \beta_1 \mathbf{1} + \mathbf{X}_2 \beta_2 + \mathbf{u}$, where $\mathbf{1}$ is an $N \times 1$ vector of ones, \mathbf{X}_2 is $N \times (k - 1)$. Assume $\mathbf{u} \sim \mathcal{N}[\mathbf{0}, \sigma^2 \mathbf{I}]$.

By subtracting means of regressors, an equivalent transformed model is

$$\mathbf{y} = \alpha_1 \mathbf{1} + \mathbf{X}_2^* \beta_2 + \mathbf{u}, \text{ where } \alpha_1 = \beta_1 + \bar{\mathbf{x}}_2' \beta_2,$$

where $\mathbf{X}_2^* = \mathbf{X}_2 - \mathbf{1}\bar{\mathbf{x}}_2'$, with $\bar{\mathbf{x}}_2' = \frac{1}{N} \mathbf{1}' \mathbf{X}_2$ a $1 \times (k - 1)$ vector of means of the regressors.

(a) Write this transformed model as

$$\mathbf{y} = \mathbf{Z}\boldsymbol{\gamma} + \mathbf{u},$$

for appropriate \mathbf{Z} and $\boldsymbol{\gamma}$.

(b) Show that $\mathbf{X}_2^* \mathbf{1} = \mathbf{0}$.

(c) Use $\hat{\boldsymbol{\gamma}} = (\mathbf{Z}'\mathbf{Z})^{-1} \mathbf{Z}'\mathbf{y}$ and part (b), even if you could not prove it, to show that OLS regression in the transformed model yields, after some algebra,

$$\begin{bmatrix} \hat{\alpha}_1 \\ \hat{\boldsymbol{\beta}}_2 \end{bmatrix} = \begin{bmatrix} \bar{y} \\ (\mathbf{X}_2^* \mathbf{X}_2^*)^{-1} \mathbf{X}_2^* \mathbf{y} \end{bmatrix}$$

(d) And show that the usual OLS estimate of the variance matrix simplifies to

$$\widehat{V} \begin{bmatrix} \widehat{\alpha}_1 \\ \widehat{\beta}_2 \end{bmatrix} = s^2 \begin{bmatrix} \frac{1}{N} & \mathbf{0} \\ \mathbf{0} & (\mathbf{X}_2^* \mathbf{X}_2^*)^{-1} \end{bmatrix}.$$

4. Now consider prediction of $E[y|\mathbf{x}_2 = \mathbf{x}_{2f}]$ using the transformed model of question 3, where you can use the results in 3(c) and 3(d). We use predicted value

$$\widehat{y}_f = \widehat{\alpha}_1 + \mathbf{x}_{2f}^* \widehat{\beta}_2,$$

where \mathbf{x}_{2f}^* is a $(k-1) \times 1$ vector with j^{th} entry $(x_{jf} - \bar{x}_j)$.

(a) Obtain $\widehat{V}[\widehat{y}_f]$. [Hint: first write \widehat{y}_f in the form $\mathbf{c}'\widehat{\gamma}$].

(b) Hence obtain a 95 percent confidence interval for \widehat{y}_f as a prediction of $E[y|\mathbf{x}_2 = \mathbf{x}_{2f}]$.

(c) For what value of \mathbf{x}_{2f} is this confidence interval the narrowest? Explain your answer.

5. Consider the OLS estimator $\widehat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$ in the multiple regression model

$$\mathbf{y} = \mathbf{X}\beta + \mathbf{u},$$

where \mathbf{y} is an $N \times 1$ vector, \mathbf{X} is an $N \times k$ matrix, β is a $k \times 1$ vector and \mathbf{u} is an $N \times 1$ vector. We assume that \mathbf{X} is fixed and $\mathbf{u} \sim N[\mathbf{0}, \Sigma]$ where $\Sigma \neq \sigma^2\mathbf{I}$.

(a) Obtain $E[\widehat{\beta}]$.

(b) Obtain $V[\widehat{\beta}]$.

(c) Give the distribution of $\widehat{\beta}$.

6. Consider the attached Stata output that uses data from the study by Joni Hersch, "Compensating Wage Differentials for Gender-Specific Job Injury Risks," American Economic Review, 1998, pp. 598-607. The original data are from the 1994 Current Population Survey, for a sample of males age 18-65.

Dependent Variable:

lnwage = Natural logarithm of hourly wage rate in dollars

Regressors:

jirate = job injury rate per 100 workers per year for the industry that the individual works in

exper = years of work experience

expersq = experience-squared

educ = years of schooling

union = 1 if union member and 0 otherwise

nonwhite = 1 if nonwhite and 0 otherwise

- (a) Which regressors are statistically significant at level 0.05 using a two-sided test? Explain.
- (b) Are the regressors jointly statistically significant at level 0.05? Explain.
- (c) State how you would set up a test of the null hypothesis that `exper` and `expersq` are jointly statistically significant at level 0.05. [There is not enough information to complete the test].
- (d) Test at level 0.05 the claim that one more year of education is associated with a 10 percent rise in wage. [Note: $\ln y/dx = (dy/y)/dx$.] State clearly details of the test and your conclusion.

```
. summarize lnwage jirate exper expersq educ union nonwhite
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lnwage	182	2.476015	.4783792	1.15268	4.454347
jirate	182	4.23074	1.529187	1.195951	7.602587
exper	182	20.38736	11.73709	0	55
expersq	182	552.647	571.8225	0	3025
educ	182	12.21703	2.704791	1	17
union	182	.1868132	.3908367	0	1
nonwhite	182	.0769231	.2672044	0	1

```
.
```

```
. regress lnwage jirate exper expersq educ union nonwhite
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Source	SS	df	MS	Number of obs =	182
Model	13.2466557	6	2.20777595	F(6, 175) =	13.71
Residual	28.174595	175	.160997685	Prob > F =	0.0000
				R-squared =	0.3198
				Adj R-squared =	0.2965
Total	41.4212507	181	.228846689	Root MSE =	.40125

lnwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
jirate	-.0173072	.0198123	-0.87	0.384	-.0564089 .0217945
exper	.0250062	.009082	2.75	0.007	.0070818 .0429307
expersq	-.0002364	.0001918	-1.23	0.219	-.0006149 .0001421
educ	.0836888	.0124466	6.72	0.000	.0591241 .1082536
union	.1123168	.0778764	1.44	0.151	-.0413811 .2660147
nonwhite	-.2754669	.1119563	-2.46	0.015	-.4964253 -.0545085
_cons	1.147848	.2028857	5.66	0.000	.74743 1.548265